

The Effect of Chronic Exercise Training on Leptin: A Systematic Review and Meta-Analysis of Randomized Controlled Trials

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Abstract

Background Leptin is a hormone associated with satiety, lipid oxidation, energy expenditure, and energy homeostasis. To date, the current body of research examining the effect of chronic exercise training on leptin has yielded inconsistent results.

Objective The purpose of this meta-analysis was to provide a quantitative estimate of the magnitude of change in leptin levels following participation in exercise interventions lasting ≥ 2 weeks.

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Methods All studies included were peer-reviewed and published in English. To be included, studies randomized human participants to an exercise training group or non-exercise comparison group for an exercise training intervention. Leptin levels were measured at baseline, during, and/or after completion of the exercise training program. Random-effects models were used to aggregate a mean effect size (ES) and 95% confidence intervals (CIs), and identify potential moderators.

Results Seventy-two randomized controlled trials met the inclusion criteria and resulted in 107 effects ($n = 3826$). The mean ES of 0.24 (95% CI 0.16–0.32, $p < 0.0001$) indicated a decrease in leptin following an exercise training program. A decrease in %Fat ($\beta = -0.07$, $p < 0.01$) was associated with a decrease in leptin after accounting for the type of control group ($\beta = -0.38$, $p < 0.0001$) used in each study.

Conclusion These results suggest that engaging in chronic exercise training (≥ 2 weeks) is associated with a decrease in leptin levels for individuals regardless of age and sex. However, a greater decrease in leptin occurred with a decreased percentage of body fat.

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Key Points

Research indicates that leptin plays a key role in the regulation of body weight, however the research regarding the effect of chronic exercise training on leptin has yielded inconsistent results.

Our results indicate that a small decrease in leptin occurs following chronic exercise training.

Observed decreases in leptin were associated with a decrease in adiposity or weight loss following chronic exercise training, after accounting for differences in study design.

1 Introduction

The prevalence of obesity has increased over recent decades, posing a significant public health problem as a number of unfavorable health consequences are associated with increased adiposity [1–4]. Identifying more effective weight loss intervention strategies to alleviate this financial and public health burden is paramount. Exercise training is used by 90% of individuals who attempt weight loss [5], however structured exercise training programs typically lead to only modest reductions in body weight [6]. A portion of the variation in weight loss following a structured exercise training intervention may be explained by hormonal control of energy intake and energy expenditure [7].

Leptin is one of the major hormones responsible for controlling energy balance and body weight by altering energy intake and energy expenditure [8, 9]. It is released in response to acute energy availability during periods of intermittent fasting, caloric restriction, and overfeeding [9–11]. Experimental studies have also indicated that the decrease in leptin observed during periods of caloric restriction is associated with increased feelings of hunger, greater desire to eat, and larger prospective food consumption [12, 13]. Furthermore, exogenous leptin administration appears to increase energy expenditure in obese and non-obese participants following weight loss [14, 15]. Longer-term clinical trials utilizing leptin administration produced a dose–response effect on weight- and fat-loss, such that greater doses of leptin yielded a greater decrease in weight and fat mass [16]. Taken together, decreased leptin levels appear to signal the body to restore energy balance by increasing appetite and decreasing energy expenditure, whereas increased leptin levels suppress appetite and increase energy expenditure [17]. However, obesity may

also be associated with leptin resistance, where individuals exhibit increased energy intake and lower levels of physical activity despite higher circulating leptin levels [18].

While it stands to reason that exercise and physical activity could potentially disrupt energy balance and impact circulating leptin levels, the chronic effects of exercise training on leptin levels are somewhat inconsistent [19]. Higher- and lower-intensity aerobic exercise decreased leptin during a 12-week intervention in obese women [20]. In addition, lower-, moderate-, and higher-intensity resistance exercise decreased leptin during a 12-month intervention in older adults [21]. In contrast, 12 weeks of aerobic exercise [22], 12 weeks of resistance training [23], or a shorter 3-week combined aerobic and resistance training yielded no change in leptin levels [24]. Finally, chronic exercise training programs have shown little effect on plasma leptin concentrations in the absence of weight loss or decreased adiposity [25–27].

Leptin has been shown to decrease following a single acute bout of aerobic exercise [28, 29] and following a single acute bout of resistance training [30]. Non-controlled experimental studies have also indicated that exercise can lower leptin in the absence of changes in body composition [31]. The acute decrease in leptin following a single bout of exercise may signal the body to increase energy intake to keep energy balance within a narrow physiological range. This compensatory increase in energy intake could potentially erase the energy deficit caused by exercise and provides a mechanism that explains why weight loss is ineffective for many individuals. However, if chronic exercise training can alter one of the key hormones regulating energy intake and energy expenditure by decreasing leptin and improve leptin resistance, perhaps specific exercise prescription can be used to target this hormone and improve the effectiveness of exercise as a treatment for obesity. Furthermore, the independent effect of chronic exercise training on leptin is still incompletely characterized, and may be influenced by a number of potential moderating factors, including study length, duration, and intensity of the exercise bout, the mode of exercise prescribed, and sex-related differences, in addition to changes in adiposity [25, 27]. As such, the primary aims of this study are to quantify the effect of chronic exercise training (≥ 2 weeks) on plasma leptin values, and identify potential moderators to determine the source of heterogeneity among effects.

2 Methods

2.1 Search Strategy

The review was conducted in accordance with the Preferred Reporting Items for Systematic Reviews and Meta-

analyses (PRISMA) statement guidelines [32]. Articles published before 14 April 2016 were located by conducting searches of the Physical Education Index ($n = 149$), PubMed ($n = 355$), Scopus ($n = 225$), SPORTDiscus ($n = 54$), and Web of Science ($n = 314$) online databases using combinations of the terms ‘exercise’, ‘training’, ‘exercise training’, ‘leptin’, ‘randomized’, ‘randomized controlled study’, and ‘randomized controlled trial’. Duplicate publications were removed and preliminary screening of the title and abstract was performed by one of the authors (MVF) for eligibility.

2.2 Study Selection

The current analysis was limited to (1) peer-reviewed publications; (2) available in English; (3) involving human subjects; (4) randomized to an exercise training or non-exercise comparison; (5) providing measures of leptin at baseline, during, and/or after completion of the exercise training program; and (6) with the length of the exercise training intervention ≥ 2 weeks. Excluded records (1) were non-peer reviewed; (2) provided a review, meta-analysis, position statement, or proposed study design; (3) utilized a cross-sectional or prospective study design; (4) included exercise as part of a multicomponent treatment (e.g. exercise plus dietary intervention), from which an effect size (ES) for exercise alone could not be calculated; or (5) compared exercise only with an active treatment (e.g. nutritional intervention or another mode of exercise). As age was to be explored as a potential moderator, no age limit was used in the current study, and retrieved articles included children/adolescents and adults. In addition, the search was not restricted to a specific range of publication dates. Potential full-text manuscripts were reviewed by three authors (MVF, EDH, CLW-R) during the data extraction process to ensure eligibility based on the criteria provided above, and resolved by consensus. A total of 133 full-text records were identified during the initial electronic database search and evaluated for inclusion in the current analysis. A flowchart of the study selection process is provided in Fig. 1. Two authors (MVF, TDW) manually reviewed reference lists from retrieved articles for additional publications not retrieved using the database search.

Manual searches revealed two additional publications [22, 33]. A request for missing data was sent to nine corresponding authors [34–42], three of whom provided the missing information or an additional publication with the requested data [35, 36, 43].

2.3 Effect Size (ES) Calculation

ESs were calculated by subtracting the mean change in the comparison condition from the mean change in the exercise

condition and dividing the difference by the pooled standard deviation (SD) of the baseline scores, and adjusted for small sample bias [44]. The inverse variance weight was used to calculate the overall mean ES. When group mean values were presented without a measure of variability [45, 46], the SD from the largest study available with similar demographic characteristics was used to estimate the ES [47]. The sample size for the intervention and control groups, mean leptin values at baseline, and during the intervention, as well as baseline measures of variability, participant age, sex, baseline body mass index (BMI), baseline %Fat, baseline fitness level, exercise mode, exercise session duration, exercise session intensity, exercise session frequency, exercise program length, year of publication, and outcome measure were independently extracted by the authors (MVF, EDH, CLW-R), with discrepancies resolved prior to aggregating effects. A decrease in leptin levels following exercise training resulted in a positive ES.

2.4 Bias Assessment

The Jadad scale was used to assess study quality and was coded by two independent reviewers (CLW-R and EDH). Discrepancies were resolved by consensus with a third reviewer (MVF) if needed. Randomization methods, blinding, and description of withdrawals were evaluated by reviewers, and scored out of 5 total points. Higher scores indicate a greater number of key criteria having been reported, and trials with scores ≥ 3 were considered of high methodological quality. In addition, bias was assessed using a funnel plot and sensitivity analysis, fail-safe N, and Egger’s test [48, 49].

2.5 Statistical Analysis

Random-effects models were used to aggregate a mean ES and 95% confidence intervals (CIs) for leptin, and to test variation in effects according to moderator variables using macros (MeanES, MetaF, and MetaReg) in IBM SPSS 23.0 (IBM Corporation, Armonk, NY, USA). Initial comparisons were made using MetaF and MetaReg based on a number of independent variables chosen a priori due to their influence on leptin levels. The definition of subgroups and moderator variables can be found in electronic supplementary Table 1.

Two-way (effects \times raters) intraclass correlation coefficients for absolute agreement were calculated to examine interrater reliability for calculated effects. The initial intraclass correlation coefficients were > 0.87 . Intraclass correlation increased to 100% after adjusting for discrepancies between reviewers.

Missing values not obtained from the literature were imputed using the mean of the available values. Multilevel

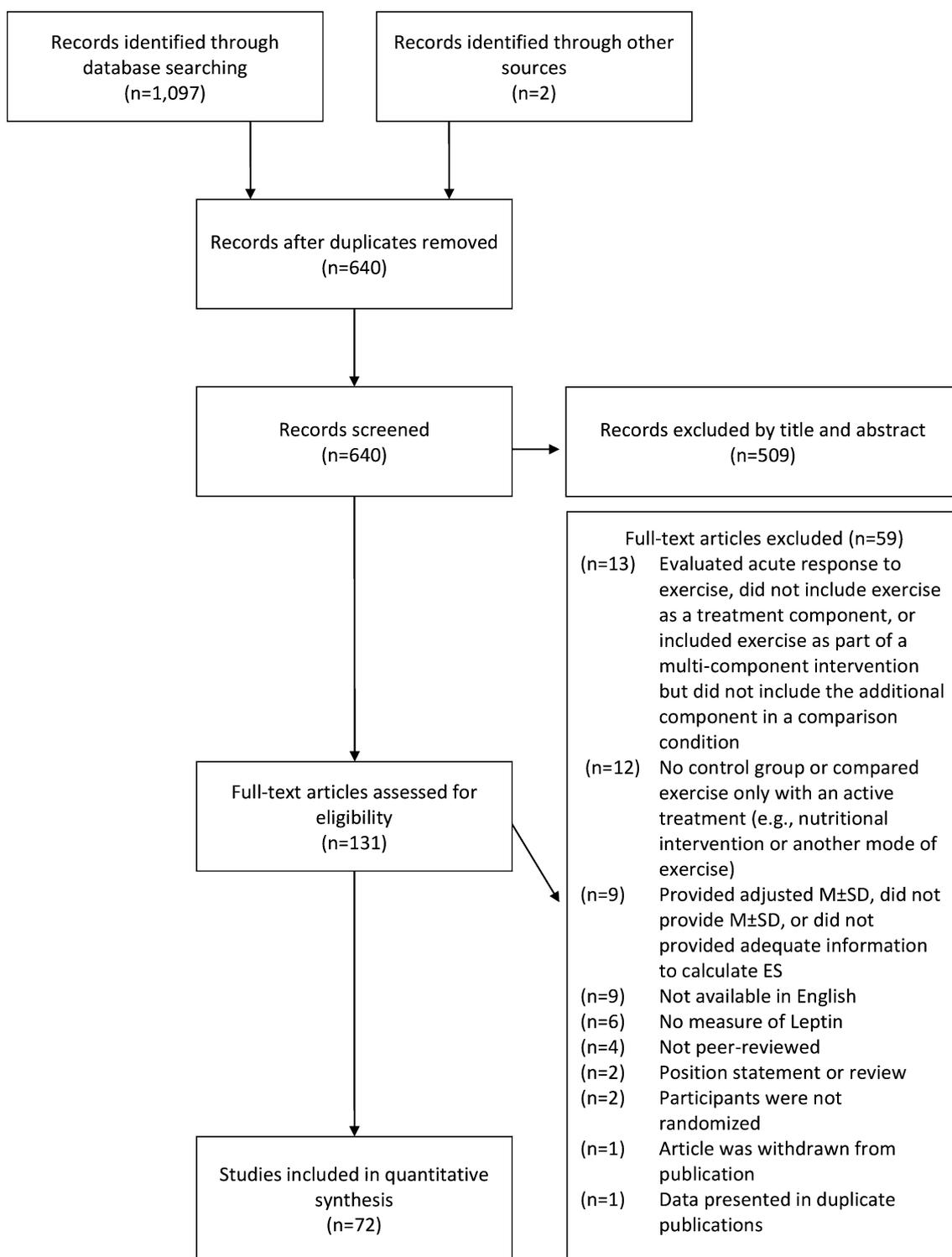


Fig. 1 Study selection process. *ES* effect size, *n* number of studies, *m* mean, *SD* standard deviation

linear regression with maximum likelihood estimation was used according to standard procedures to adjust for between-study variance and the correlation between effects nested within studies [50, 51]. This was required as multiple effects were gathered from studies with multiple

intervention groups or reporting subgroup comparisons in addition to the mean change. The data analysis for the multilevel model was performed using SAS 9.4 (SAS Institute Inc., Cary, NC, USA). Data available for study

Table 1 Summary of subgroup analysis

Effect moderator	Effects (<i>k</i>)	Mean ES	95% CI	<i>p</i> value	Between-group comparison <i>p</i> value
Age group, years					
< 20	6	-0.1743	-0.5098, 0.1611	0.3083	
20–55	63	0.2889	0.1844, 0.3935	<0.0001	
> 55	38	0.2293	0.1140, 0.3446	0.0001	0.0344
Sex					
Female	49	0.2846	0.1824, 0.3869	<0.0001	
Male	38	0.3253	0.1947, 0.4559	<0.0001	
Mixed	20	-0.0288	-0.1933, 0.1357	0.7315	0.0018
Baseline weight status					
Normal weight	23	0.2528	0.0692, 0.4365	0.0070	
Overweight	39	0.2647	0.1312, 0.3983	0.0001	
Obese	45	0.2242	0.1085, 0.3400	0.0001	0.8988
Exercise mode					
Aerobic	55	0.2589	0.1568, 0.3610	<0.0001	
Resistance	21	0.3446	0.1413, 0.5480	0.0009	
Combined	31	0.1541	0.0071, 0.3011	0.0399	0.2927
Exercise intensity					
Low	38	0.1714	0.0383, 0.3045	0.0116	
Moderate	25	0.3120	0.1454, 0.4786	0.0002	
Vigorous	44	0.2637	0.1460, 0.3814	<0.0001	0.3887
Clinical diagnosis					
Healthy participants	19	0.3595	0.1630, 0.5560	0.0003	
Type 1 diabetes mellitus	1	0.2692	-0.7715, 1.3099	0.6122	
Type 2 diabetes mellitus	13	0.0839	-0.1046, 0.2724	0.3830	
Overweight/obese	58	0.2817	0.1959, 0.3674	<0.0001	
Chronic kidney disease	1	1.0781	0.2253, 1.9309	0.0132	
Polycystic ovarian syndrome	2	0.8035	-0.6260, 0.7931	0.8175	
Pregnancy	1	-0.7979	-1.3445, -0.2513	0.0052	
Cardiovascular disease	1	0.3000	-0.2844, 0.8844	0.3144	
Breast cancer	1	0.1155	-0.3872, 0.6182	0.6614	
Colorectal cancer	1	0.0887	-0.7712, 0.9486	0.8398	
Esophageal cancer	1	0.3352	-0.4313, 1.1017	0.3914	
Knee osteoarthritis	1	-0.2560	-0.7026, 0.1906	0.2612	
Metabolic syndrome	6	0.1375	-0.0916, 0.3666	0.2394	
Spinal cord injury	1	0.7601	-0.2748, 1.7950	0.1500	0.0055

CI confidence interval, *ES* effect size

and participant characteristics are presented as mean \pm SD unless otherwise indicated.

3 Results

3.1 Mean ES

Effects analyzed included data collected from 3826 participants following exercise training. These individuals

received some form of exercise training in an intervention setting that ranged from 4 to 166 participants (29.8 ± 31.7) per effect. The age of participants ranged from 8.2 to 71.2 years (47.4 ± 15.4 years), and participants were predominantly female ($67.6 \pm 42.5\%$), with 18.7% of effects obtained from samples of mixed sex. Eighty-seven effects were obtained from sex-specific samples (35.5% male only, 45.8% female only). Average BMI and %Fat across all effects were 30.0 ± 3.0 kg/m² and $38.0 \pm 8.4\%$ Fat, respectively.

Training interventions varied in design but utilized aerobic training only (51.4%), resistance training only (19.6%), or a combination of training modes (29.0%). Exercise interventions were 3 weeks up to 24 months in length (7.5 ± 4.8 months) and trained participants from 2 to 7 days (4.0 ± 1.0 days), reaching 24–90 min (52.3 ± 13.7 min) of exercise per session at the conclusion of the study. Studies included in the current analysis are presented in electronic supplementary Table 2.

The cumulative results from 107 effects extracted from 72 articles, published between 1998 and 2016, indicate that chronic exercise training can effectively reduce leptin levels (ES 0.2430, 95% CI 0.1647–0.3212, $p < 0.0001$) (Fig. 2). The estimate of the effect increased slightly after accounting for the nesting of multiple effects within a single study (ES 0.2437, 95% CI 0.1390–0.3483, $p < 0.0001$).

3.2 Homogeneity of Results

The significant decrease in leptin following chronic exercise training presented moderate heterogeneity ($Q_{106} = 194.8$, $p < 0.001$, $I^2 = 45.6\%$), with sampling error accounting for 71.7% of the observed variance [44, 52]. Based on a significant Q statistic and an I^2 near 50% indicating moderate heterogeneity, the variability in leptin following chronic exercise training is greater than would have occurred naturally based on study sample error. The null hypothesis for homogenous distribution was rejected, and subgroup and moderator analyses were performed to determine the source of variability among effects.

3.3 Subgroup and Moderator Analysis

Subgroup comparisons indicated that the change in leptin among older adults (> 55 years) was no different ($k = 38$, ES 0.23, 95% CI 0.11–0.34, $p < 0.001$) than the change in younger adults (20–55 years) ($k = 63$, ES 0.29, 95% CI 0.18–0.39, $p < 0.001$), however both appeared to be greater than children/adolescents (< 20 years) ($k = 6$, ES -0.17 , 95% CI -0.51 to 0.16, $p = 0.3083$) as significant differences were noted between groups ($p = 0.0344$). The effect of chronic exercise training on leptin was significant for both men ($k = 38$, ES 0.31, 95% CI 0.20–0.42, $p < 0.0001$) and women ($k = 49$, ES 0.27, 95% CI 0.19–0.35, $p < 0.001$), with no sex differences observed ($p = 0.5606$). Finally, exercise training interventions that utilized a no-treatment control group ($k = 78$, ES 0.33, 95% CI 0.24–0.42, $p < 0.001$) yielded larger effects compared with training interventions that used an ‘active’ control group (i.e. diet vs. diet + exercise) ($k = 29$, ES 0.05, 95% CI -0.08 to 0.19, $p = 0.4653$), with significant differences observed between conditions ($p = 0.0008$). Additional subgroup comparisons are presented in Table 1.

The independent effects of each potential moderating variable were examined using univariate meta-regression and are presented in Table 2. In the multivariate model, control group comparison ($\beta = -0.38$, standard error [SE] 0.0901, $p < 0.001$) and change in %Fat ($\beta = -0.0684$, SE 0.03, $p = 0.0078$) were associated with the effect of exercise on leptin, collectively accounting for 14.4% of the variation in effects. These data indicate that a greater decrease in leptin was observed in individuals who experienced a reduction in %Fat, after accounting for the type of control group. Both parameters, control group comparison ($\beta = -0.33$, SE 0.09, $p = 0.0012$) and change in %Fat ($\beta = -0.08$, SE 0.03, $p = 0.0092$) remained

Table 2 Summary of univariate moderator analysis

Effect moderator	Unstandardized β	95% CI	R^2 (%)	p value
Age	-0.0016	-0.0064, 0.0032	0.36	0.5228
Percentage female	-0.0873	-0.2653, 0.0907	0.80	0.3366
Baseline BMI	-0.0098	-0.0333, 0.0137	0.57	0.4148
Change BMI ES	0.1197	-0.0092, 0.2303	3.82	0.0338
Baseline %Fat	-0.0042	-0.0134, 0.0049	0.71	0.3673
Change %Fat ES	-0.0178	-0.0657, 0.0302	0.45	0.4685
Exercise session duration	-0.0037	-0.0094, 0.0019	1.41	0.1976
Exercise program length	-0.0093	-0.0257, 0.0072	1.06	0.2685
Exercise frequency	-0.0182	-0.0588, 0.0952	0.19	0.6433
Exercise intensity	0.0436	-0.0454, 0.1325	0.79	0.3373
Comparison condition	-0.2791	-0.4430, -0.1153	9.20	0.0008

BMI body mass index, CI confidence interval, ES effect size, %Fat relative adiposity

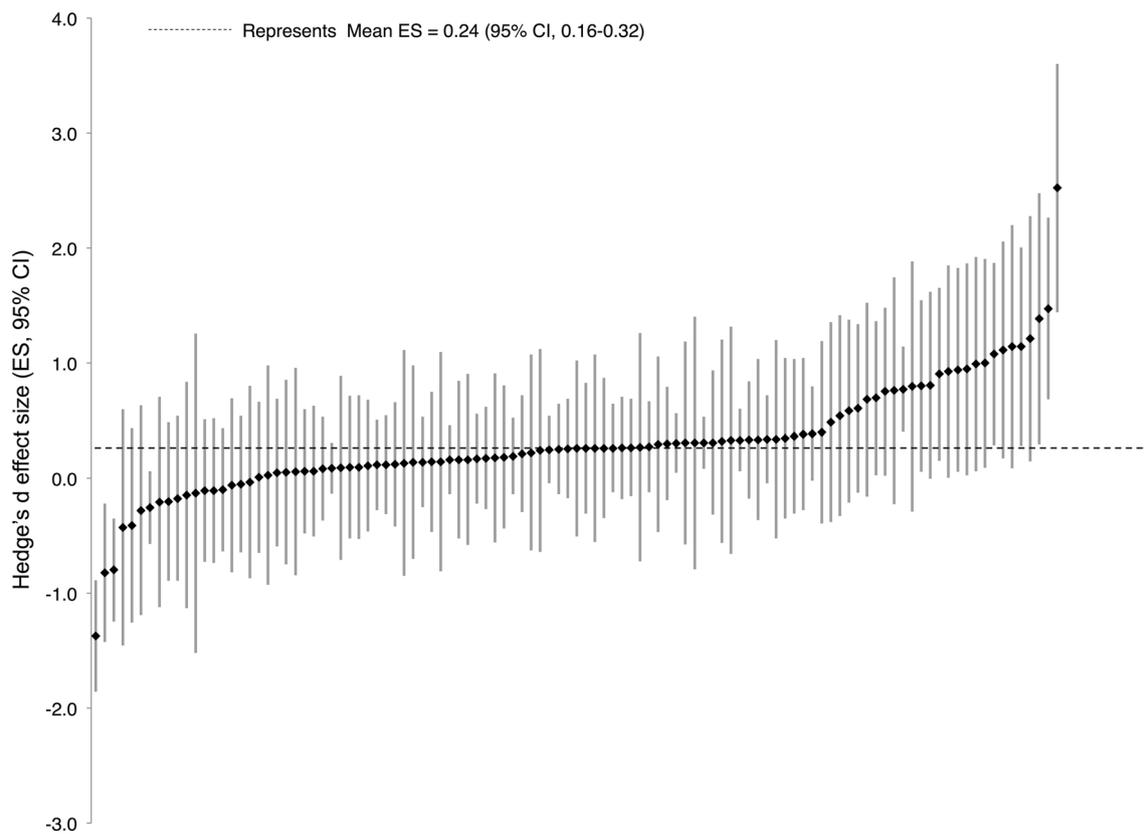


Fig. 2 Forest plot of Hedges' *d* effect sizes. The aggregated Hedges' *d* is the random-effects mean effect size for chronic exercise training on leptin weighted by the inverse variance. Data are presented as effect size and 95% CI. *CI* confidence interval, *ES* effect size

significantly associated with the change in leptin after adjusting for the nesting of effects within studies.

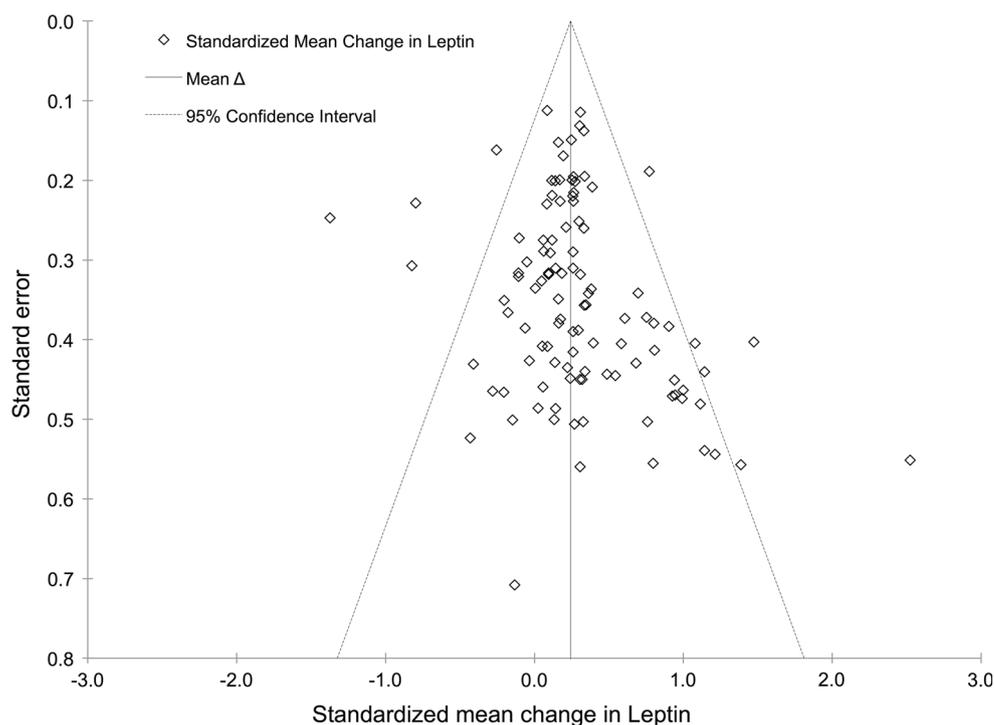
3.4 Assessment of Bias

Twenty-nine effects (27.1%) were obtained from high-quality studies, with overall quality of scores for effects in the current analysis ranging from 1 to 4 (1.9 ± 0.9). All 107 effects (100%) were randomized, with 29 effects (27.1%) gathered from studies that provided a description of the randomization process used and that the current authors deemed appropriate (table of random numbers, computer-generated order, etc.). A complete description of participants who were enrolled in the study but did not complete the observation period or who were not included in the analysis was provided for 53 effects (49.5%). Although it is often not possible to 'blind' participants to an exercise treatment condition, 12 effects (11.2%) described the procedures used to blind researchers during data collection and analysis.

The number of unpublished or unretrieved null effects needed to overturn the significant results was estimated

using a fail-safe N^+ [53]. The fail-safe N^+ for the effect of exercise training on leptin using a random-effects model collapsed to a fixed-effects model after one iteration, and yielded $N^+ = 1708.4$ effects. The robust fail-safe N^+ indicated that publication bias can be 'safely ignored' [48]. A funnel plot was created by scattering the treatment ES exercise training on leptin against their SE [54], as recommended when using a standardized mean difference ES [54–56]. These data are presented in Fig. 3 and indicate low likelihood of potential bias. In addition, 10 of the 107 effects (9.3%) outside of the 95% CI were identified using the funnel plot to identify potential outliers. Sensitivity analysis removing these 10 effects had minimal effect on our estimate of the effect of exercise on leptin (ES 0.2408, 95% CI 0.1854–0.2962, $p < 0.001$) for the remaining 97 effects, but reduced the observed heterogeneity in effects ($Q_{96} = 63.7201$, $p = 0.9955$, $I^2 = 0.00\%$). Finally, potential publication bias was addressed using Egger's test [49]. Further supporting the fail-safe N^+ , funnel plot, and sensitivity analysis, Egger's test indicated that the effect of exercise training on leptin was not subject to bias [$F(1, 105) = 0.196$, $p = 0.659$].

Fig. 3 Funnel plot of Hedges' d effect size versus study standard error. The aggregated Hedges' d is the random-effects mean effect size for chronic exercise training on leptin plotted against the standard error



4 Discussion

The cumulative evidence gathered from peer-reviewed research published between 1998 and 2016 indicates that chronic exercise training (≥ 2 weeks) is associated with a decrease in plasma leptin. Our results provide an estimate of the magnitude of the effect of chronic exercise training on leptin levels, and indicate that a decrease in leptin occurs as a result of chronic exercise training, regardless of age and sex. Aerobic, resistance, and concurrent exercise training all yielded a significant decrease in leptin, with no differences noted between the modes of exercise. In addition, exercise intensity, duration, frequency, and program length were not associated with a change in leptin. Although a decrease in %Fat appears to be one of the most significant factors associated with a decrease in leptin levels, chronic exercise training appears to elicit an independent and significant decrease in leptin levels regardless of change in adiposity.

Dietary interventions have consistently led to a decrease in circulating leptin levels and appear to yield similar decreases in leptin levels when compared with exercise training [57]. While either may be effective stand-alone treatment options, combined diet and exercise training treatments have been associated with the largest decrease in leptin levels, especially when accompanied by weight loss and decrease in %Fat [22, 43, 58]. The physiologic impact of this decrease in leptin on weight loss success during diet and exercise interventions remains unclear. Previous research has indicated that while higher leptin levels may decrease feelings of hunger and increase energy

expenditure in normal-weight individuals, obese individuals appear to be leptin-resistant, where higher circulating leptin levels may not be able to induce these same responses [9]. The acute decrease in leptin concentration in response to an energy deficit may serve as a protective mechanism to control energy expenditure during periods of low energy availability [59]. During an exercise program, these mechanisms may counteract a weight loss program in order to keep total daily energy expenditure within a narrow physiologic range. These protective mechanisms may be partly responsible for weight regain as previously obese individuals tend to have lower leptin levels when compared with their normal weight, 'never obese' peers [60]. On the other hand, the decrease in leptin following chronic exercise training observed in the current study may be due to improved leptin sensitivity [61–63], and may represent the body establishing a new physiologic 'set point'. As this analysis was limited to exercise interventions longer than 2 weeks, the improved leptin sensitivity may enhance further weight loss during periods of negative energy balance as %Fat explained only a portion of the variation in the change in leptin following exercise training [64].

Combined dietary and exercise interventions show promise when treating obesity and may compliment pharmaceutical interventions [65, 66]. When elevated plasma leptin levels do not effectively reduce energy intake, this often leads to further weight gain, and ultimately increases leptin resistance [67]. Although somewhat speculative, previous researchers have argued that improved leptin sensitivity would result in weight loss almost exclusively

due to decreased adiposity [68] as exogenous leptin administration produced a decrease in fat mass that explained over 95% of weight loss, during the 24-week study period, in obese participants [16]. The increased leptin sensitivity of a combined treatment program including diet, exercise, and pharmaceutical treatment could potentially induce greater weight loss than any treatment independently, and would have tremendous clinical significance when coupled with the improvements in fat-free mass that occur as a result of exercise training. However, due to the exploratory nature of this analysis, an attempt should be made to explore these hypotheses in future experimental studies.

The result of the subgroup analysis indicated that the effect of chronic exercise training varied by the clinical population being studied, however only two of the clinical populations studied (pregnancy and knee osteoarthritis) presented negative mean ES. In addition, 10 of the 14 clinical populations studied yielded fewer than three effects, and group ES ranged from -0.7979 to 0.8035 . The largest number of effects came from three distinct populations (healthy, overweight/obese, and type 2 diabetes mellitus participants). Evidence supporting the hypothesis that the effect of chronic exercise training varies by clinical population is somewhat inconclusive despite the observed, and perhaps spurious, significant between-group differences. Rather, these results should highlight the need for further research examining whether the presence of a specific disease or condition alters the efficacy of exercise training.

Based on the current results, improvements in body composition were the strongest predictors of the observed mean effect. Determining the best treatment strategy for overweight/obese patients or participants is paramount. As no single component of the exercise prescription was associated with the observed decrease in leptin, practitioners are encouraged to design their physical activity and exercise training interventions based on the current guidelines for weight loss and weight management, and accumulate at least 250 min of moderate-intensity physical activity per week [69]. The current study also indicates that aerobic and resistance training may yield a similar decrease in leptin, which aligns with previous evidence that has suggested both modes of exercise can yield similar weight loss, especially when accompanied by a restricted-calorie dietary program [70]. Perhaps when working with individuals to achieve weight loss, health-fitness professionals should emphasize a holistic approach to human movement combining structured exercise training with increased lifestyle-based physical activity. Taken together, physical activity is generally better than none at all, more physical activity is often better than less, and including higher-volume or higher-intensity activities is appropriate when under the guidance of a trained professional [71, 72].

4.1 Limitations

Because of the nature of meta-analysis and reviews of secondary data, the current review is not without limitations. First, the current study calculated a standardized mean effect of exercise training using summary data from published peer-reviewed manuscripts. Using raw individual participant data rather than summary descriptive statistics may have provided added precision for the current analysis, however this approach was not feasible due to the large number of publications and wide range of years included in our electronic search [73]. Second, the current study was limited to peer-review publications and excluded ‘grey literature’ from conference abstracts, dissertations, and other non-peer-reviewed sources. Although grey literature was not included in order to limit the current analysis to studies of high methodological quality, excluding these other sources has been shown to have little impact on the overall estimated ES [74–76]. As with all systematic reviews, it is difficult to estimate the extent to which unpublished, unidentified, and unretrieved studies may have potentially influenced the results of the current study [53]. The authors feel confident that this review identified the vast majority of the relevant publications as 97.3% of the publications were identified using the electronic database search, as well as identifying additional articles using a manual search of references [77–80]. It should be noted that fewer than half of the studies in the current review included a complete description of participant randomization, attrition, and missingness (46.9%). Improved transparency when reporting data management and analysis should be a priority in future studies.

Specific to the current analysis, the authors made the decision to include children/adolescents in the current review a priori because it was unclear how much research had examined the effect of chronic exercise training on leptin within this age group. In the end, only six effects were collected from articles with an average age < 20 years. A sensitivity analysis removing children/adolescents did not impact the overall mean effect (ES 0.2430 , 95% CI 0.1647 – 0.3212 , $p < 0.0001$ in the original manuscript vs. ES 0.2558 , 95% CI 0.1841 – 0.3275 , $p < 0.0001$ when excluding children/adolescents). In addition, excluding children had no impact on the results of the moderator analysis comparing ‘younger’ and ‘older’ adults (0.2715 , 95% CI 0.1777 – 0.3653 vs. 0.2296 , 95% CI 0.1310 – 0.3281 , respectively; $p = 0.5457$). The magnitude of decrease in leptin following chronic exercise training appear to be much larger in ‘younger’ and ‘older’ adults, however caution should be used when interpreting the between-group differences based on age due to the relatively small number of studies involving children/adolescents. Based on the available data, it is unclear if these apparent differences

represent the true effect of chronic exercise training on leptin in the population or represent a statistical artifact due to the paucity of research in children. As such, we recommend that future high-quality original research should further examine the effect of chronic exercise training on leptin in children/adolescents.

5 Conclusions

Based on the cumulative results from peer-reviewed research published between 1998 and 2016, chronic exercise training (≥ 2 weeks) produces a small but consistent decrease in plasma leptin levels regardless of age and sex. The effect of exercise training appears to be largely driven by improvements in body composition, specifically a decrease in %Fat. However, the decrease also appears to suggest increased leptin sensitivity, even in the absence of weight loss or decrease in %Fat. These results further support the role of exercise training as part of an effective weight management program.

Author Contributions MVF conceptualized and designed the study, coded and analyzed effects, carried out the initial analysis, drafted the initial manuscript, and approved the final manuscript as submitted. EDH coded and analyzed effects, reviewed and revised the initial manuscript, and approved the final manuscript as submitted. CLW-R coded and analyzed effects, reviewed and revised the initial manuscript, and approved the final manuscript as submitted. TDW coded and analyzed effects, reviewed and revised the initial manuscript, and approved the final manuscript as submitted. WCD coded and analyzed effects, reviewed and revised the initial manuscript, and approved the final manuscript as submitted.

Compliance with Ethical Standards

Data Availability Statement Data used for these analyses are available in a public repository through the University of Alabama, which does not issue datasets with DOIs (non-mandated deposition). The data can be downloaded directly from <http://ir.ua.edu/handle/123456789/3480> in SPSS or Microsoft Excel file format.

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